

Lecture 23: Central Limit Theorem

Statistics 251

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Lecture Outline

Theorem Statement

Moment Generating Function: Part II

Proof of the theorem

Examples

Where are we?

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Revision

Recall the weak law of large numbers,

Let X_1, X_2, \dots be a sequence of independent and identically distributed random variables, each having finite mean $E[X_i] = \mu$.

Then, for any $\varepsilon > 0$,

$$P \left\{ \left| \frac{X_1 + \dots + X_n}{n} - \mu \right| \geq \varepsilon \right\} \rightarrow 0 \quad \text{as} \quad n \rightarrow \infty$$

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In some sense means, $\left| \frac{X_1 + \dots + X_n - n\mu}{n} \right| \rightarrow 0$. What about

$$\frac{X_1 + \dots + X_n - n\mu}{\sqrt{n}}$$

or in other words, how fast does it converge?

Central limit theorem

Let X_1, X_2, \dots be a sequence of independent and identically distributed random variables, each having mean μ and variance σ^2 . Then the distribution of

$$\frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}$$

tends to the standard normal as $n \rightarrow \infty$. That is, for $-\infty < a < \infty$,

$$\mathbb{P} \left\{ \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}} \leq a \right\} \rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^a e^{-x^2/2} dx \text{ as } n \rightarrow \infty$$

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Definition:

$$M(t) = E \left[e^{tX} \right]$$
$$= \begin{cases} \sum_x e^{tx} p(x) & \text{if } X \text{ is discrete with mass function } p(x) \\ \int_{-\infty}^{\infty} e^{tx} f(x) dx & \text{if } X \text{ is continuous with density } f(x) \end{cases}$$

When calculating moments

$$M^n(0) = E [X^n] \quad n \geq 1$$

Moment Generating Function of sum of independent random variables

If X and Y are independent and have moment generating functions $M_X(t)$ and $M_Y(t)$, respectively. Then $M_{X+Y}(t)$, is given by

$$\begin{aligned}M_{X+Y}(t) &= E \left[e^{t(X+Y)} \right] \\&= E \left[e^{tX} e^{tY} \right] \\&= E \left[e^{tX} \right] E \left[e^{tY} \right] \\&= M_X(t) M_Y(t).\end{aligned}$$

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If $\{X_i\}_{i=1}^n$ are identical independent random variables from the same distribution with moment generating function M_X , what is $M_{\sum_{i=1}^n X_i}$?

Lemma: moment generating function decides the distribution

- ▶ Let Z_1, Z_2, \dots be a sequence of random variables having distribution functions F_{Z_n} and moment generating functions M_{Z_n} , $n \geq 1$.
- ▶ Let Z be a random variable having distribution function F_Z and moment generating function M_Z .
- ▶ If $M_{Z_n}(t) \rightarrow M_Z(t)$ for all t , then $F_{Z_n}(t) \rightarrow F_Z(t)$ for all t at which $F_Z(t)$ is continuous.

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We first assume mean $\mu = 0$, variance $\sigma^2 = 1$. (Why we can do this?)

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$$\left[M\left(\frac{t}{\sqrt{n}}\right) \right]^n.$$

Let $L(t) = \log M(t)$, then

$$\log \left(\left[M\left(\frac{t}{\sqrt{n}}\right) \right]^n \right) = tnL(t/\sqrt{n})$$

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Proof conti

What is the limit of $nL(t/\sqrt{n})$ when $n \rightarrow \infty$?

$$\begin{aligned}\lim_{n \rightarrow \infty} \frac{L(t/\sqrt{n})}{n^{-1}} &= \lim_{n \rightarrow \infty} \frac{-L'(t/\sqrt{n})n^{-3/2}t}{-2n^{-2}} \quad (\text{by L'H\^opital's rule}) \\ &= \lim_{n \rightarrow \infty} \left[\frac{L'(t/\sqrt{n})t}{2n^{-1/2}} \right] \\ &= \lim_{n \rightarrow \infty} \left[\frac{-L''(t/\sqrt{n})n^{-3/2}t^2}{-2n^{-3/2}} \right] \quad (\text{by L'H\^opital's rule}) \\ &= \lim_{n \rightarrow \infty} \left[L''\left(\frac{t}{\sqrt{n}}\right) \frac{t^2}{2} \right] \\ &= \frac{t^2}{2} \quad (\text{as } L''(0) = \frac{M(0)M''(0) - [M'(0)]^2}{[M(0)]^2} = 1)\end{aligned}$$

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CLT only applies when $N \rightarrow \infty$. Empirically, only when $N \geq 30$, you can have confidence to apply it. If $N < 10$, unless otherwise noted, we should use Markov inequalities.

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When the random variable converges to normal distribution, in some case we need to know the exact value of integral

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$$\phi(a) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^a e^{-x^2/2}.$$

a	-2.58	-1.96	-1.645	0	1.645	1.96	2.58
$\phi(a)$	0.005	0.025	0.05	0.5	0.95	0.975	0.995

Bound for sum of rolling dice

Let $X_i, i = 1, \dots, 10$, be independent random variables, each uniformly distributed over $(0, 1)$. Calculate an approximation to

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First,

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Then

$$\begin{aligned} P\{29.5 \leq X \leq 40.5\} &= P \left\{ \frac{29.5 - 35}{\sqrt{\frac{350}{12}}} \leq \frac{X - 35}{\sqrt{\frac{350}{12}}} \leq \frac{40.5 - 35}{\sqrt{\frac{350}{12}}} \right\} \\ &\approx 2\Phi(1.0184) - 1 \\ &\approx .692 \end{aligned}$$